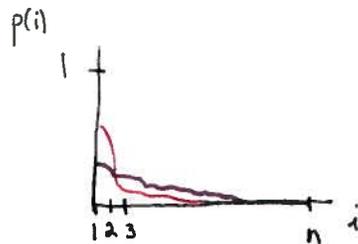


Lecture 15:

Testing monotonicity of distributions

Testing & Learning Monotone Distributions (over totally ordered domain)

Def. p over $[n]$ is "monotone decreasing"
if $\forall i \in [n-1] \quad p(i) \geq p(i+1)$



Monotonicity Tester:

- if p monotone increasing, Pass with prob $\geq 3/4$
- if p ϵ -far in L_1 dist from mon increasing, Fail with prob $\geq 3/4$

Useful tool: "Birge Decomposition"

(note: this is a different decomposition than in homework (upcoming)
in particular, it is oblivious!)

decompose domain $1..n$ into $l = \Theta\left(\frac{\log \epsilon n}{\epsilon}\right) \approx \Theta\left(\frac{\log n}{\epsilon}\right)$ intervals

$$I_1^\epsilon, I_2^\epsilon, \dots, I_l^\epsilon \quad \text{s.t.}$$

$$|I_{kn}^\epsilon| = \lfloor (1+\epsilon)^k \rfloor$$

← will drop ϵ
in notation
once it's fixed

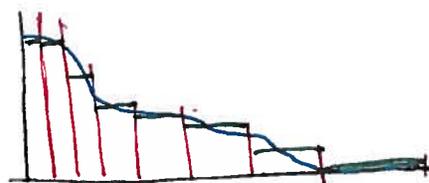
$$|I_1^\epsilon| = |I_2^\epsilon| = \dots = 1$$

$$|I_a^\epsilon| = |I_{an}^\epsilon| = \dots = 2$$

but then at some point the sizes grow
exponentially

define "flattened distribution"

$$\forall 1 \leq j \leq l \quad \forall i \in I_j \quad \tilde{q}(i) = \frac{q(I_j)}{|I_j|}$$



← assign all elements in same interval the same probability

note: $q(I_j) = \tilde{q}(I_j)$

Birge's Thm if q mon decreasing then $\|\tilde{q} - q\|_1 < \epsilon$

Coroll if q ϵ -close to mon decreasing then $\|\tilde{q} - q\|_1 < O(\epsilon)$

Testing Algorithm:

Take samples of q
do uniformity test for each partition (using samples that fell in it):
(if not enough samples then pass) "FAIL"

$w_j \leftarrow$ # samples that fell in partition j
use LP to verify w close to monotone
(that is dist which gives wt w_j to each Birge bucket + uniform within each bucket)

if $> \epsilon$ fraction of weight is in partitions that fail.

note this is LP on $O(\log n)$ vars

how can we do this? \tilde{q} isn't even if q monotone, exactly uniform. See problem from next hw set.

How many samples?

for each partition with enough weight, say $\frac{\epsilon}{\log n}$, need $\frac{\sqrt{n}}{\epsilon^2}$ samples
 $\approx O(\sqrt{n} \text{ polylog } n \cdot \text{poly } \frac{1}{\epsilon})$
need $\frac{\sqrt{n} \cdot \log n}{\epsilon^3}$ for each one
need another $\log \log n$ for union bound

(note: this can be improved!!)

Last step:

difficulty

sampling error might make w_j 's look non monotone



purple is not monotone
but is close

good thing: only $\frac{\log n}{\epsilon}$ variables!

can be solved via brute force
LP (actually quite efficient)

⋮

so: monotone \neq likely to pass

ϵ -far from monotone p : either (1) non uniform in buckets
or (2) w far from monotone

} more details
on next
page

Slightly changing perspective...

What if we know dist q is monotone, can we learn it?

Yes! use sampling to estimate $\tilde{q}_\epsilon(I_j)$'s

Birge's Thm \Rightarrow Can learn monotone distributions to w/in ϵ L_1 error
in $\Theta(\frac{1}{\epsilon^3} \log n)$ samples.

Correctness: (high level idea)

q monotone:

• Birge $\Rightarrow \|q - \tilde{q}\| < \epsilon' < \epsilon$

• Claim: (ignoring partitions with max wt $\frac{\epsilon}{2h}$)
 for $\leq \frac{\log n}{\epsilon}$ partitions, the $\frac{\min \text{ value of } q}{\max \text{ value of } q} \leq \epsilon'$

An issue: total # of "bad" partitions is small, but also need total weight to be small or need to fix algorithm

"Bad partitions" are $< \epsilon$ fraction of all partitions

• for "good" partitions, uniformity test likely to pass [see HW in future]

• also, w is close to \tilde{q} which is monotone, so w is close to monotone

q ϵ -far from monotone:

• assume q "likely to pass"

\Rightarrow most Birge buckets close to uniform

w close to monotone

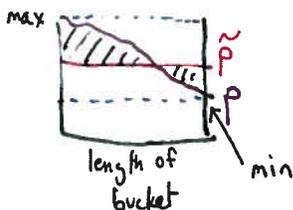
} correct to monotone \tilde{q} which is uniform on buckets s.t. average bucket wt. is monotone

$|q - \tilde{q}|$ is small $\Rightarrow q$ close to monotone
 + \tilde{q} close to monotone

thus ϵ -far \Rightarrow likely to fail

Proof of Birge's Thm :

Error in bucket



gross upper bound on error:
 $\leq (\max - \min) \cdot \text{bucket length}$

Partition of Intervals:

- Size 1 Intervals $|I_j| = 1$
- Short Intervals $|I_j| < \frac{1}{\epsilon}$
- Long Intervals $|I_j| \geq \frac{1}{\epsilon}$

no error on these!
 ← if we have ^{any} short intervals, there are $\Omega(\frac{1}{\epsilon})$ of these
 if not, we can learn the distribution

↔ if we have these then
 max prob $\leq \epsilon$ (since # size 1 intervals is $\Omega(\frac{1}{\epsilon})$)

total error $\leq \sum_{j=1}^l |I_j| \cdot (\max \text{ prob in } I_j - \min \text{ prob in } I_j)$

$= \underbrace{\sum_{\text{size 1 intervals}} 1 \cdot 0}_0$ + $\underbrace{\sum_{\text{short intervals}} |I_j| (\max - \min)}$ + $\underbrace{\sum_{\text{long intervals}} |I_j| (\max - \min)}$

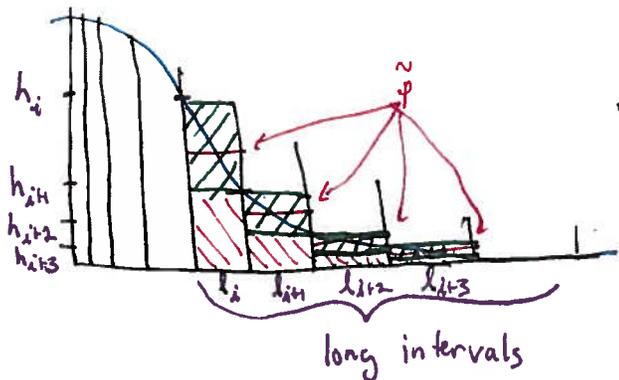
0 since no difference

omitted: idea is bound similarly to the long intervals but need to group together intervals of same size

see below

↑ therefore min size 1 interval has prob $\leq \epsilon$ which upper bounds later probabilities too since p is monoton

Picture for long intervals:



green rectangles = upper bound on error

error $\leq (h_i - h_{i+1}) \cdot l_i + (h_{i+1} - h_{i+2}) l_{i+1} + (h_{i+2} - h_{i+3}) l_{i+2} + \dots$

$= h_i l_i + h_{i+1} (l_{i+1} - l_i) + h_{i+2} (l_{i+2} - l_{i+1}) + h_{i+3} (l_{i+3} - l_{i+2})$

all h_i 's in this area are $< \epsilon$!

positive, $+ \approx \epsilon \cdot l_{i+1}$ by way that we partitioned

$\leq \epsilon [l_i + \sum h_i l_{i+1}]$

get rid of this when bounding short intervals

this is area of red rectangles, which is upper bounded by p so sum is ≤ 1