

Lecture 16:

Hypothesis Testing

A useful tool: Hypothesis Testing

Given collection of distributions \mathcal{H} , at least one has high accuracy for describing $p \leftarrow$ given via samples cut out one of collection that is close to p .

How many samples in terms of $|\mathcal{H}| + \text{domain size?}$

Why is this different than testing closeness uniformly?
Do we have the same lower bounds?

NO

Since p is guaranteed to be close to some $q \in \mathcal{H}$, all bets are off!!

A "subtool": allows comparing two hypothesis

Thm Given sample access to P
Given h_1, h_2 hypothesis distributions (fully known to algorithm)
Given accuracy parameter ϵ' , confidence δ'
Algorithm "Choose" takes $O(\log(1/\delta')/(\epsilon')^2)$ samples + outputs

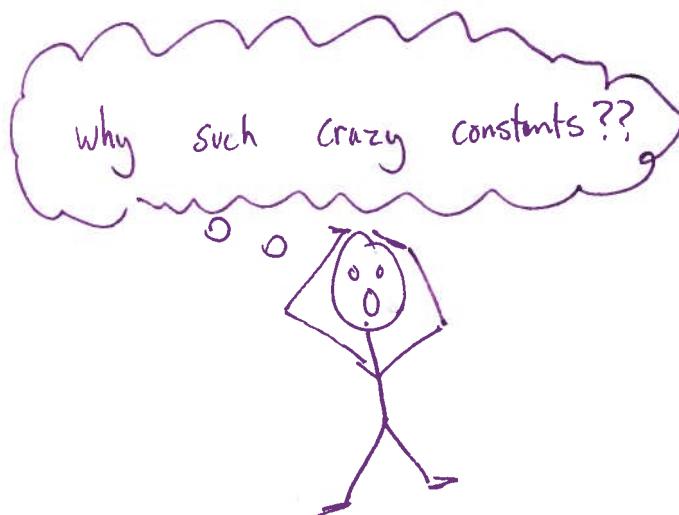
- If both h_1, h_2 far, no guarantees
- If one is close, you output something pretty close. } $h \in \{h_1, h_2\}$. If one of h_1, h_2 has $\|h_i - p\| < \epsilon'$ then with prob $\geq 1 - \delta'$, output h_j has $\|h_j - p\| < 12\epsilon'$

Actually, will prove something stronger:

Thm p given via samples
 h_1, h_2 fully known + p is ϵ' -close to \checkmark at least one of h_1, h_2
 ϵ', δ' given

Algorithm "Choose" takes $O(\log(1/\delta') (1/\epsilon')^2)$ samples
+ outputs $h \in \{h_1, h_2\}$ satisfying:

- (1) if h_i more than $12\epsilon'$ -far from p , unlikely to output it as $\underbrace{\text{very bad}}_{2e^{-m\epsilon'^2/2}}$ $\underbrace{\text{winner or tie}}$
- (2) if h_i more than $10\epsilon'$ -far, unlikely to output as winner
 $\underbrace{\text{not that bad}}$ \uparrow
 might tie
 but won't win



Proof of "Subtool":

idea: wlog h_1 is ϵ' -close,
 if h_2 is $10\epsilon'$ -close, then either output ok as "winner" or "tie"
 else, if h_2 is not $10\epsilon'$ -close but is $12\epsilon'$ -close, ok to "tie" or output h_1
 else, h_2 is $12\epsilon'$ -far from h_1 + $11\epsilon'$ -far from p
 so samples from p will fall in
 "difference" between $h_1 + h_2$ + will output h_1

Algorithm Choose: Input p, h_1, h_2 samples explicit description

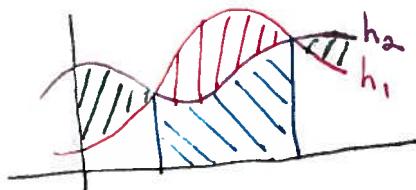
First same definitions:

$$A = \{x \mid h_1(x) > h_2(x)\}$$

$$a_1 = h_1(A)$$

$$a_2 = h_2(A)$$

$$\text{note } \|h_1 - h_2\|_1 = 2(a_1 - a_2)$$



green area = red area
 L_1 dist = green + red
 red area = $a_1 - a_2$

blue area = a_2
 blue + red area = a_1

this is important!

→ 1. if $a_1 - a_2 \leq 5\epsilon'$ declare "tie" + return h_1
 (no samples needed)

additive factor in constants

2. draw $m = 2 \cdot \frac{\log \frac{1}{\delta'}}{(\epsilon')^2}$ samples $s_1 \dots s_m$ from p

$$3. \alpha \leftarrow \frac{1}{m} |\{i \mid s_i \in A\}|$$

if $p = h_1, E[\alpha] = a_1$
 if $p = h_2, E[\alpha] = a_2$

4. if $\alpha > a_1 - \frac{3}{2}\epsilon'$ return h_1

else if $\alpha < a_2 + \frac{3}{2}\epsilon'$ return h_2

else declare "tie" + return h_1

another additive factor in constants
 (will see breakdown on next page)

Why does it work?

$$E[\alpha] = p(A)$$

- if reach step 2, whp (via Chernoff) $|\alpha - E[\alpha]| \leq \frac{\varepsilon'}{2}$

if $\|p - h_1\|_1 > 12\varepsilon'$ then since other is $\leq \varepsilon'$ distance,
 or $\|p - h_2\|_1 > 12\varepsilon'$ $\|h_1 - h_2\|_1 > 11\varepsilon'$

so will reach step 2

if p ε' -close to h_1 , whp $\alpha > a_1 - \varepsilon' - \frac{\varepsilon'}{2}$
 ↑
 from closeness to h_1 ↗ Sampling error
 so output h_1 ↘ affects constants

else, p is $12\varepsilon'$ far from h_1
 but ε' -close to h_2

$$\text{whp } \alpha < a_2 + \varepsilon' + \frac{\varepsilon'}{2}$$

if one of h_1, h_2 ε' -close from $\dots \Rightarrow$ return h_2 whp
 and other is $\geq 10\varepsilon'$ far but not $12\varepsilon'$ far \dots

if $a_1 - a_2 \leq 5\varepsilon'$ then declares draw, so neither are declared "winner"

else $\|h_1 - h_2\|_1 > 9\varepsilon'$ far

+ similar reasoning shows that
 medium far will not win (in fact, will lose)

recall:

$$\|h_1 - h_2\|_1 = 2(a_1 - a_2)$$

if both are $10\varepsilon'$ -close, might output h_1, h_2 or "tie"

The Cover Method

a method for learning distributions

def C is a ϵ -cover of D if $\forall p \in D$
 $\exists g \in C$ s.t. $\|p-g\|_1 \leq \epsilon$

\uparrow set of distributions
 \uparrow set of distributions (big)

\uparrow set of distributions (smaller)

Why useful?
 hopefully C is much smaller than D - allows us to "approx"
 note C not unique

Big improvement Thm \exists algorithm, given $p \in D$, which takes
 union bnd over size of C $\Rightarrow O(\frac{1}{\epsilon^2} \log |C|)$ samples of p + outputs $h \in C^D$
 s.t. $\|h-p\|_1 \leq 6\epsilon$ with prob $\geq 9/10$
not D !!

Pf.

since $p \in D$, $\exists g \in C^D$ s.t. $\|p-g\|_1 \leq \delta$
 (but there could be more than 1) \leftarrow we just need to find one, not even required to return p

will run Choose on p with every pair $g_1, g_2 \in C^D$
 if g doesn't win all of its "matches" then it loses
 to someone that is not so bad

Furthermore can show that w.h.p. there is a g' s.t.
 g' wins or ties all matches. (best g never loses, any one that ties it is $\leq 10\epsilon$ far)
 need all matches to give correct output - union bound on $\binom{|C|}{2}$ matches

The cover method

Example 1: learning distribution of a coin

$$\text{domain} = \{0, 1\}$$

need to learn bias

$$\text{Here } \mathcal{C} = \mathbb{R}$$

$$\text{if use } \mathcal{C} = \left\{ 0, \frac{1}{k}, \frac{2}{k}, \dots, \frac{k-1}{k}, 1 \right\}$$

$$\text{then } \forall \text{ bias } p, \text{ let } \frac{i}{k} \leq p \leq \frac{i+1}{k}$$

$$\text{then picking } \tilde{p} = \frac{i}{k} \text{ gives } \|p - \tilde{p}\|_1 = \left| \frac{i}{k} - p \right| + \left| \left(\frac{i+1}{k} \right) - p \right| \leq \frac{2}{k}$$

$$\text{so using } k = \Theta(\frac{1}{\epsilon}) \text{ gives } \|p - \tilde{p}\|_1 \leq \epsilon$$

$$|\mathcal{C}| = k+1 = \Theta(\frac{1}{\epsilon}), \text{ #samples needed by cover method is } O(\frac{1}{\epsilon^2} \cdot \log \frac{1}{\epsilon})$$

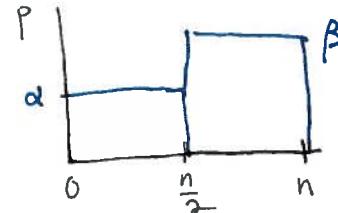
Example 2: 2-bucket distributions

now need to specify α and β

$$\text{so } \mathcal{C} = \left\{ \left(\frac{i}{k}, \frac{j}{k} \right) \mid i, j \in \{0, \dots, k\} \right\}$$

$$|\mathcal{C}| = \Theta(\frac{1}{\epsilon})^2$$

$$\# \text{samples is } O(\frac{1}{\epsilon^2} \log \frac{1}{\epsilon})$$



Example 3: monotone distributions

$$\text{Birge} \Rightarrow \mathcal{C} = \left\{ \left(\frac{i_1}{k}, \dots, \frac{i_{\lceil \log n / \epsilon \rceil}}{k} \right) \mid i_1, i_2, \dots \in \{0, \dots, k\} \right\}$$

$$|\mathcal{C}| = \Theta\left(\frac{1}{\epsilon} (\log n)^{\lceil \log n / \epsilon \rceil}\right) \Rightarrow \# \text{samples is } O\left(\frac{1}{\epsilon^3} \log n \cdot \log \frac{1}{\epsilon}\right)$$