

# Probabilistic Graphical Models

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## *Reminder of lecture 2*

- An alternative representation for joint distributions is as an **undirected graphical model** (also known as **Markov random fields**)
- As in BNs, we have one node for each random variable
- Rather than CPDs, we specify (non-negative) **potential functions** over sets of variables associated with cliques  $C$  of the graph,

$$p(x_1, \dots, x_n) = \frac{1}{Z} \prod_{c \in C} \phi_c(\mathbf{x}_c)$$

$Z$  is the **partition function** and normalizes the distribution:

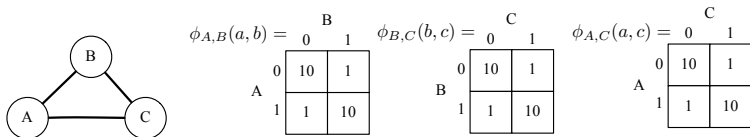
$$Z = \sum_{\hat{x}_1, \dots, \hat{x}_n} \prod_{c \in C} \phi_c(\hat{\mathbf{x}}_c)$$

# Undirected graphical models

$$p(x_1, \dots, x_n) = \frac{1}{Z} \prod_{c \in C} \phi_c(\mathbf{x}_c),$$

$$Z = \sum_{\hat{x}_1, \dots, \hat{x}_n} \prod_{c \in C} \phi_c(\hat{\mathbf{x}}_c)$$

Simple example (potential function on each edge encourages the variables to take the same value):



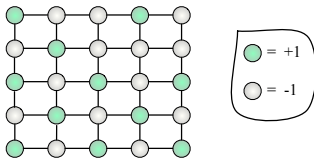
$$p(a, b, c) = \frac{1}{Z} \phi_{A,B}(a, b) \cdot \phi_{B,C}(b, c) \cdot \phi_{A,C}(a, c),$$

where

$$Z = \sum_{\hat{a}, \hat{b}, \hat{c} \in \{0,1\}^3} \phi_{A,B}(\hat{a}, \hat{b}) \cdot \phi_{B,C}(\hat{b}, \hat{c}) \cdot \phi_{A,C}(\hat{a}, \hat{c}) = 2 \cdot 1000 + 6 \cdot 10 = 2060.$$

## Example: Ising model

- Theoretical model of interacting atoms, studied in statistical physics and material science
- Each atom  $X_i \in \{-1, +1\}$ , whose value is the direction of the atom spin
- The spin of an atom is biased by the spins of atoms nearby on the material:



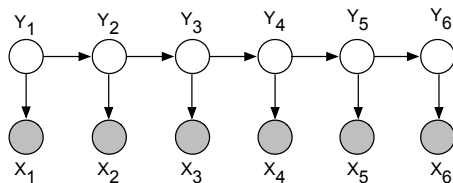
$$p(x_1, \dots, x_n) = \frac{1}{Z} \exp \left( \sum_{i < j} w_{i,j} x_i x_j - \sum_i u_i x_i \right)$$

- When  $w_{i,j} > 0$ , nearby atoms encouraged to have the same spin (called **ferromagnetic**), whereas  $w_{i,j} < 0$  encourages  $X_i \neq X_j$
- Node potentials  $\exp(-u_i x_i)$  encode the bias of the individual atoms
- Scaling the parameters makes the distribution more or less spiky

- Markov random fields
  - 1 Bayesian networks  $\Rightarrow$  Markov random fields (*moralization*)
  - 2 Hammersley-Clifford theorem (conditional independence  $\Rightarrow$  joint distribution factorization)
- Conditional models
  - 3 Discriminative versus generative classifiers
  - 4 Conditional random fields

# Converting BNs to Markov networks

What is the equivalent Markov network for a hidden Markov model?



# Moralization of Bayesian networks

- Procedure for converting a Bayesian network into a Markov network
- The **moral graph**  $\mathcal{M}[G]$  of a BN  $G = (V, E)$  is an undirected graph over  $V$  that contains an undirected edge between  $X_i$  and  $X_j$  if
  - 1 there is a directed edge between them (in either direction)
  - 2  $X_i$  and  $X_j$  are both parents of the same node



(term historically arose from the idea of “marrying the parents” of the node)

- The addition of the moralizing edges leads to the loss of some independence information, e.g.,  $A \rightarrow C \leftarrow B$ , where  $A \perp B$  is lost

# Converting BNs to Markov networks

- 1 Moralize the directed graph to obtain the undirected graphical model:



- 2 Introduce one potential function for each CPD:

$$\phi_i(x_i, \mathbf{x}_{pa(i)}) = p(x_i | \mathbf{x}_{pa(i)})$$

- So, converting a hidden Markov model to a Markov network is simple:





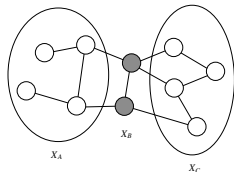
# Factorization implies conditional independencies

- $p(\mathbf{x})$  is a *Gibbs distribution* over  $G$  if it can be written as

$$p(x_1, \dots, x_n) = \frac{1}{Z} \prod_{c \in C} \phi_c(\mathbf{x}_c),$$

where the variables in each potential  $c \in C$  form a clique in  $G$

- Recall that conditional independence is given by graph separation:



- Theorem (**soundness of separation**): If  $p(\mathbf{x})$  is a Gibbs distribution for  $G$ , then  $G$  is an I-map for  $p(\mathbf{x})$ , i.e.  $I(G) \subseteq I(p)$

*Proof:* Suppose  $\mathbf{B}$  separates  $\mathbf{A}$  from  $\mathbf{C}$ . Then we can write

$$p(\mathbf{X}_A, \mathbf{X}_B, \mathbf{X}_C) = \frac{1}{Z} f(\mathbf{X}_A, \mathbf{X}_B) g(\mathbf{X}_B, \mathbf{X}_C).$$

# Conditional independencies implies factorization

- Theorem (**soundness of separation**): If  $p(\mathbf{x})$  is a Gibbs distribution for  $G$ , then  $G$  is an I-map for  $p(\mathbf{x})$ , i.e.  $I(G) \subseteq I(p)$
- What about the converse? We need one more assumption:
- A distribution is **positive** if  $p(\mathbf{x}) > 0$  for all  $\mathbf{x}$
- Theorem (**Hammersley-Clifford**, 1971): If  $p(\mathbf{x})$  is a positive distribution and  $G$  is an I-map for  $p(\mathbf{x})$ , then  $p(\mathbf{x})$  is a Gibbs distribution that factorizes over  $G$
- Proof is in book (as is counter-example for when  $p(\mathbf{x})$  is not positive)
- This is important for **learning**:
  - Prior knowledge is often in the form of conditional independencies (i.e., a graph structure  $G$ )
  - Hammersley-Clifford tells us that it suffices to search over Gibbs distributions for  $G$  – allows us to *parameterize* the distribution

- Markov random fields
  - ① Bayesian networks  $\Rightarrow$  Markov random fields (*moralization*)
  - ② Hammersley-Clifford theorem (conditional independence  $\Rightarrow$  joint distribution factorization)
- Conditional models
  - ③ Discriminative versus generative classifiers
  - ④ Conditional random fields

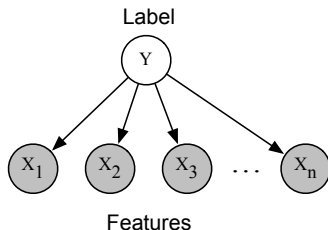
- There is often significant flexibility in choosing the structure and parameterization of a graphical model

**It is important to understand the trade-offs**

- In the next few slides, we will study this question in the context of e-mail classification

# From lecture 1... naive Bayes for classification

- Classify e-mails as spam ( $Y = 1$ ) or not spam ( $Y = 0$ )
  - Let  $1 : n$  index the words in our vocabulary (e.g., English)
  - $X_i = 1$  if word  $i$  appears in an e-mail, and 0 otherwise
  - E-mails are drawn according to some distribution  $p(Y, X_1, \dots, X_n)$
- Words are conditionally independent given  $Y$ :

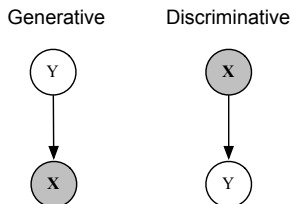


- Prediction given by:

$$p(Y = 1 \mid x_1, \dots, x_n) = \frac{p(Y = 1) \prod_{i=1}^n p(x_i \mid Y = 1)}{\sum_{y \in \{0,1\}} p(Y = y) \prod_{i=1}^n p(x_i \mid Y = y)}$$

# Discriminative versus generative models

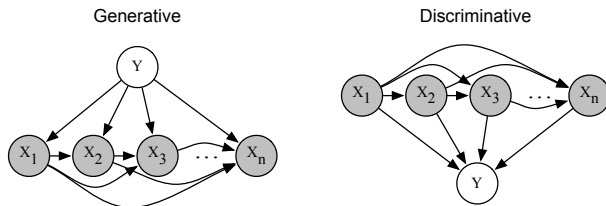
- Recall that these are **equivalent** models of  $p(Y, \mathbf{X})$ :



- However, suppose all we need for prediction is  $p(Y | \mathbf{X})$
- In the left model, we need to estimate *both*  $p(Y)$  and  $p(\mathbf{X} | Y)$
- In the right model, it suffices to estimate just the **conditional distribution**  $p(Y | \mathbf{X})$ 
  - We never need to estimate  $p(\mathbf{X})!$
  - Not possible to use this model when  $\mathbf{X}$  is only partially observed
  - Called a **discriminative** model because it is only useful for discriminating  $Y$ 's label

# Discriminative versus generative models

- Let's go a bit deeper to understand what are the trade-offs inherent in each approach
- Since  $\mathbf{X}$  is a random vector, for  $Y \rightarrow \mathbf{X}$  to be equivalent to  $\mathbf{X} \rightarrow Y$ , we must have:



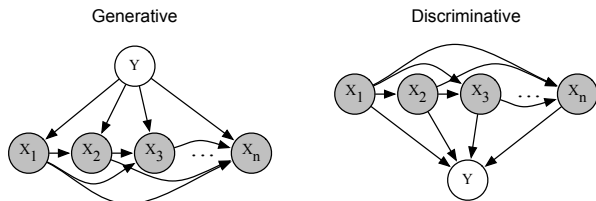
We must make the following choices:

- 1 In the generative model, how do we parameterize  $p(X_i \mid \mathbf{X}_{pa(i)}, Y)$ ?
- 2 In the discriminative model, how do we parameterize  $p(Y \mid \mathbf{X})$ ?

# Discriminative versus generative models

We must make the following choices:

- 1 In the generative model, how do we parameterize  $p(X_i | \mathbf{X}_{pa(i)}, Y)$ ?
- 2 In the discriminative model, how do we parameterize  $p(Y | \mathbf{X})$ ?



- 1 For the generative model, assume that  $X_i \perp \mathbf{X}_{-i} | Y$  (**naive Bayes**)
- 2 For the discriminative model, assume that

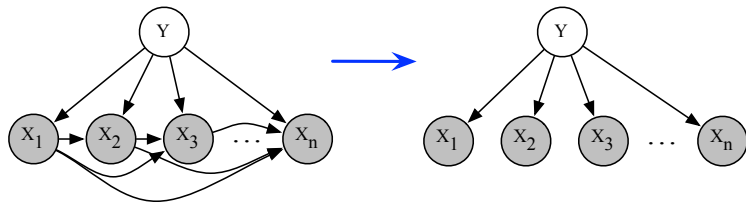
$$p(Y = 1 | \mathbf{x}; \alpha) = \frac{e^{\alpha_0 + \sum_{i=1}^n \alpha_i x_i}}{1 + e^{\alpha_0 + \sum_{i=1}^n \alpha_i x_i}} = \frac{1}{1 + e^{-\alpha_0 - \sum_{i=1}^n \alpha_i x_i}}$$

This is called **logistic regression**. (To simplify the story, we assume  $X_i \in \{0, 1\}$ )



# Naive Bayes

- 1 For the generative model, assume that  $X_i \perp \mathbf{X}_{-i} \mid Y$  (**naive Bayes**)

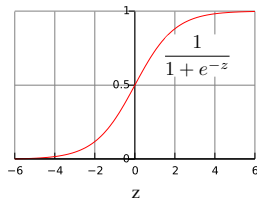
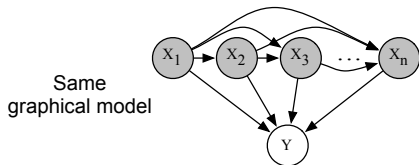


# Logistic regression

- 2 For the discriminative model, assume that

$$p(Y = 1 \mid \mathbf{x}; \alpha) = \frac{e^{\alpha_0 + \sum_{i=1}^n \alpha_i x_i}}{1 + e^{\alpha_0 + \sum_{i=1}^n \alpha_i x_i}} = \frac{1}{1 + e^{-\alpha_0 - \sum_{i=1}^n \alpha_i x_i}}$$

Let  $z(\alpha, \mathbf{x}) = \alpha_0 + \sum_{i=1}^n \alpha_i x_i$ . Then,  $p(Y = 1 \mid \mathbf{x}; \alpha) = f(z(\alpha, \mathbf{x}))$ , where  $f(z) = 1/(1 + e^{-z})$  is called the **logistic function**:



# Discriminative versus generative models

- 1 For the generative model, assume that  $X_i \perp \mathbf{X}_{-i} \mid Y$  (**naive Bayes**)
- 2 For the discriminative model, assume that

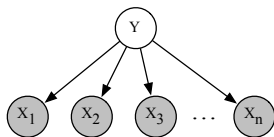
$$p(Y = 1 \mid \mathbf{x}; \alpha) = \frac{e^{\alpha_0 + \sum_{i=1}^n \alpha_i x_i}}{1 + e^{\alpha_0 + \sum_{i=1}^n \alpha_i x_i}} = \frac{1}{1 + e^{-\alpha_0 - \sum_{i=1}^n \alpha_i x_i}}$$

- In problem set 1, you showed **assumption 1**  $\Rightarrow$  **assumption 2**
- Thus, every conditional distribution that can be represented using naive Bayes can *also* be represented using the logistic model
- What can we conclude from this?

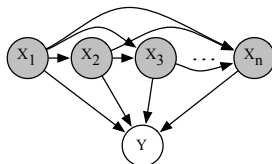
**With a large amount of training data, logistic regression will perform at least as well as naive Bayes!**

# Discriminative models are powerful

Generative (naive Bayes)



Discriminative (logistic regression)



- Logistic model does *not* assume  $X_i \perp \mathbf{X}_{-i} \mid Y$ , unlike naive Bayes
- This can make a big difference in many applications
- For example, in spam classification, let  $X_1 = 1[\text{“bank” in e-mail}]$  and  $X_2 = 1[\text{“account” in e-mail}]$
- Regardless of whether spam, these always appear together, i.e.  $X_1 = X_2$
- Learning in naive Bayes results in  $p(X_1 \mid Y) = p(X_2 \mid Y)$ . Thus, naive Bayes **double counts the evidence**
- Learning with logistic regression sets  $\alpha_i = 0$  for one of the words, in effect ignoring it (there are other equivalent solutions)

# Generative models are still very useful

- 1 Using a conditional model is only possible when  $\mathbf{X}$  is always observed
  - When some  $X_i$  variables are unobserved, the generative model allows us to compute  $p(Y | \mathbf{X}_e)$  by marginalizing over the unseen variables
- 2 Estimating the generative model using maximum likelihood is more **efficient** (statistically) than discriminative training
  - When only a small amount of training data is available, naive Bayes can outperform logistic regression
  - Relevant only when the model is reasonably accurate (i.e., the data generating distribution respects the implied independencies)
  - We will return to these questions in the second half of the course

# Conditional random fields (CRFs)

- **Conditional random fields** are undirected graphical models of conditional distributions  $p(\mathbf{Y} \mid \mathbf{X})$ 
  - $\mathbf{Y}$  is a set of **target variables**
  - $\mathbf{X}$  is a set of **observed variables**
- We typically show the graphical model using just the  $\mathbf{Y}$  variables
- Potentials are a function of  $\mathbf{X}$  and  $\mathbf{Y}$

# Formal definition

- A CRF is a Markov network on variables  $\mathbf{X} \cup \mathbf{Y}$ , which specifies the conditional distribution

$$P(\mathbf{y} \mid \mathbf{x}) = \frac{1}{Z(\mathbf{x})} \prod_{c \in \mathcal{C}} \phi_c(\mathbf{x}_c, \mathbf{y}_c)$$

with partition function

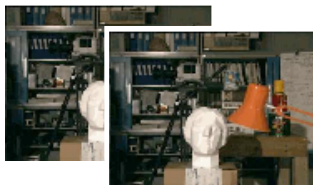
$$Z(\mathbf{x}) = \sum_{\hat{\mathbf{y}}} \prod_{c \in \mathcal{C}} \phi_c(\mathbf{x}_c, \hat{\mathbf{y}}_c).$$

- As before, two variables in the graph are connected with an undirected edge if they appear together in the scope of some factor
- The only difference with a standard Markov network is the normalization term – before marginalized over  $\mathbf{X}$  and  $\mathbf{Y}$ , now only over  $\mathbf{Y}$

# CRFs in computer vision

- Undirected graphical models very popular in applications such as computer vision: segmentation, stereo, de-noising
- Grids are particularly popular, e.g., pixels in an image with 4-connectivity

input: two images



output: disparity



- Not encoding  $p(\mathbf{X})$  is the main strength of this technique, e.g., if  $\mathbf{X}$  is the image, then we would need to encode the distribution of natural images!
- Can encode a rich set of features, without worrying about their distribution



# Parameterization of CRFs

- Factors may depend on a large number of variables
- We typically parameterize each factor as a log-linear function,

$$\phi_c(\mathbf{x}_c, \mathbf{y}_c) = \exp\{\mathbf{w} \cdot \mathbf{f}_c(\mathbf{x}_c, \mathbf{y}_c)\}$$

- $\mathbf{f}_c(\mathbf{x}_c, \mathbf{y}_c)$  is a feature vector
- $\mathbf{w}$  is a weight vector which is typically learned – we will discuss this extensively in later lectures

# NLP example: named-entity recognition

- Given a sentence, determine the people and organizations involved and the relevant locations:  
“Mrs. Green spoke today in New York. Green chairs the finance committee.”
- Entities sometimes span multiple words. Entity of a word not obvious without considering its *context*
- CRF has one variable  $X_i$  for each word, which encodes the possible labels of that word
- The targets are, for example, “B-person, I-person, B-location, I-location, B-organization, I-organization”
  - Having beginning (B) and outcome (I) allows the model to segment adjacent entities

# NLP example: named-entity recognition

This is typically represented having two factors for each word:

- $\phi_t^1(Y_t, Y_{t+1})$  represents dependencies between neighboring target variables
- $\phi_t^2(Y_t, X_1, \dots, X_T)$  represents dependencies between a target and its context in the word sequence

The graphical model looks like:

